

Example (1). Using the first step-decomposition theorem (1) given in the previous Lecture # 7. Let $\{X_n, n = 0, 1, \dots\}$ be a MC with TPM $\mathbf{M} = (p_{ij})$, on the state space $SS = \{0, 1\}$. we can easily establish the following:

$$f_{00}^{(1)} = \Pr(T_{00} = 1) = \Pr(X_1 = 0 | X_0 = 0) = p_{00},$$

$$\begin{aligned} f_{00}^{(n)} = \Pr(T_{00} = n) &= \sum_{k \in SS - \{0\}} p_{0k} f_{k0}^{(n-1)} = p_{01} f_{10}^{(n-1)} = p_{01} p_{11} f_{10}^{(n-2)} \\ &= \dots = p_{01} (p_{11})^{n-2} p_{10}, \quad n \geq 2, \end{aligned}$$

$$\text{Similarity, } f_{11}^{(n)} = \begin{cases} p_{11}, & n = 1 \\ p_{10} (p_{00})^{n-2} p_{01}, & n \geq 2 \end{cases},$$

$$f_{01}^{(n)} = (p_{00})^{n-1} p_{01}, \text{ and } f_{10}^{(n)} = (p_{11})^{n-1} p_{10}, \text{ for } n \geq 1.$$

Bernoulli Trials:

Consider the tossing of an unfair coin with success given by $\{\text{head} = s\}$ with probability p and failure given by $\{\text{tail} = f\}$ with probability q , and $p + q = 1$. The sample space is $S_1 = \{s, f\}$. If the coin is tossed twice, then the sample space $S_2 = \{ss, sf, fs, ff\}$ is an ordered set from the Cartesian product $S_1 \times S_1$. The cardinality of S_2 is $2^2 = 4$. The probability of two successes is p^2 and two failures is q^2 . If we toss the coin n times, then the resulting sample space is $S_n = S_1 \times S_1 \times \dots \times S_1$ and the cardinality of S_n is 2^n . The probability of n successes is p^n .

Def. (Bernoulli trials)

Bernoulli trials are repeated functionally independent trials with only two events s and f for each trial. The trials are also statistically independent with the two events s and f in each trial having probabilities p and $q = 1 - p$, respectively. These are called independent identically distributed (iid) trials.

Def. (Bernoulli process)

If X_n is a random variable denotes the number of successes in the trial n , The stochastic process $\{X_n, n = 1, 2, \dots\}$ is called Bernoulli process with probability of success p ($0 \leq p \leq 1$), if it satisfies:

- 1- The random variables X_1, X_2, \dots are independent
- 2- The event $\{X_n = 1\}$ denotes to the fail in the trial number n , while the event $\{X_n = 0\}$ denotes to the success in the trial n , i.e.,

$$X_n = \begin{cases} 1 & \text{for success } s \text{ in the } n^{\text{th}} \text{ trial} \\ 0 & \text{for failure } f \text{ in the } n^{\text{th}} \text{ trial} \end{cases},$$

with probabilities $\Pr(X_n = 1) = p$, and $\Pr(X_n = 0) = q = 1 - p \forall n = 1, 2, \dots$

In the Bernoulli process the parameter set is $T = \{1, 2, \dots\}$ and the state space is $SS = \{0, 1\}$ the two are discrete.

The statistics of Bernoulli process

Mean: $E[X_n] = \mu_{X_n} = 1 \cdot \Pr(X_n = 1) + 0 \cdot \Pr(X_n = 0) = p$

Second moment: $E[X_n^2] = \mu_{X_n^2} = 1^2 \cdot \Pr(X_n = 1) + 0^2 \cdot \Pr(X_n = 0) = p$

Variance: $Var(X_n) = E[X_n^2] - (E[X_n])^2 = p - p^2 = p(1 - p)$

Probability generating function

$$G_X(z) = E[z^{X_n}] = z^0 \times \Pr(X_n = 0) + z^1 \times \Pr(X_n = 1) = q + zp$$

Autocorrelation function:

$$R_x(m, n) = E[X_m X_n] = \begin{cases} 1^2 \cdot p + 0^2 \cdot (1-p) = p, & m = n \\ E[X_m]E[X_n] = p^2, & m \neq n \end{cases}$$

Auto-covariance:

$$\text{Cov}(X_m, X_n) = R_x(m, n) - E[X_m]E[X_n] = \begin{cases} p - p^2 = p(1-p), & m = n \\ p^2 - p^2 = 0, & m \neq n \end{cases}$$

Normalized auto-covariance:

$$\rho_x(m, n) = \begin{cases} 1, & m = n \\ 0, & m \neq n \end{cases}$$

Since the mean is independent of time n and the autocorrelation depends only on the time difference $m-n$, the process X_n is stationary.

Def. (Binomial process)

If $\{X_n, n=1,2,\dots\}$ represents Bernoulli process with probability of success, p ($0 \leq p \leq 1$), and N_n represents the number of successes during the first trials until the completion of the trial n :

$$N_n = \begin{cases} X_1 + X_2 + \dots + X_n, & n \geq 1 \\ 0, & n = 0 \end{cases}$$

where the increments $\{X_i\}$ form a family of independent $\{0,1\}$ -valued random variables.

The process $\{N_n, n=1,2,\dots\}$ is called a binomial process.

The statistics of the binomial process

Mean: Since the probability of each success is p , the mean value of N_n can be obtained using the linearity of the expectation operator:

$$\mu_{N_n} = E[N_n] = E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] = nE[X_i] = np$$

Second moment: The second moment of N_n can be given by

$$E[N_n^2] = E\left[\left(\sum_{i=1}^n X_i\right)^2\right] = \sum_{i=1}^n E[X_i^2] + \sum_{i \neq j} \sum_j E[X_i]E[X_j] = np + n(n-1)p^2$$

Variance:

$$\text{Var}(N_n) = E[N_n^2] - (E[N_n])^2 = np + n(n-1)p^2 - n^2p^2 = np(1-p).$$

Probability generating function

$$G_{N_n}(z) = E[z^{N_n}] = E[z^{X_1+X_2+\dots+X_n}] = E[z^{X_1}]E[z^{X_2}]\dots E[z^{X_n}] = (q + zp)^n$$

Autocorrelation:

$$\begin{aligned} R_N(m, n) &= E[N_m N_n] = E\left[\sum_{j=1}^n \sum_{i=1}^m X_i X_j\right] \\ &= E\left[\sum_{i=1}^m X_i^2\right] + \sum_{j=1}^m \sum_{\substack{i=1 \\ i \neq j}}^m E[X_i X_j] + \sum_{j=1}^{n-m} \sum_{\substack{i=1 \\ i \neq j}}^m E[X_i X_j] \\ &= mp + m(m-1)p^2 + m(n-m)p^2 = \begin{cases} mp(1-p) + mnp^2 & \text{for } m \leq n \\ np(1-p) + mnp^2 & \text{for } n \leq m \end{cases} \\ &= p(1-p)\min(m, n) + mnp^2, \quad m, n > 0. \end{aligned}$$

Auto-covariance:

$$\text{Cov}(N_m, N_n) = R_N(m, n) - E[N_m]E[N_n] = p(1-p)\min(m, n), \quad m, n > 0.$$

Binomial process is a Markov chain

If the random variable N_n denotes the number of successes during the number n of Bernoulli's trials, where the probability of success in any one trial is p , the sequence of $\{N_n, n=1, 2, \dots\}$ is a MC, the probability of transition in one step is

$$p_{ij} = \Pr(N_{n+1} = j | N_n = i) = \Pr(X_{n+1} = j - i) = \begin{cases} p, & j = i + 1, \\ 1 - p, & j = i, \\ 0, & \text{otherwise} \end{cases},$$

and the transition probabilities after n -step is given by

$$p_{ij}^{(n)} = \binom{n}{j-i} p^{j-i} q^{n-j+1}, \quad j = i, \dots, n+i.$$

If N_n is the number of successes in the first n Bernoulli trials, with probability of a success in any one trial is p , then the sequence of random variables $\{N_n, n = 1, 2, \dots\}$ is a MC, with $N_0 = 0$, since the process $\{N_n\}$ has the discrete parameter set $T = \{1, 2, \dots\}$ and the discrete state space $SS = \{0, 1, 2, \dots\}$, and satisfies the Markov property: for all $i, k, i_1, \dots, i_{n-1} \in SS$ we have

$$\begin{aligned} & \Pr(N_{n+1} = k | N_n = i, N_{n-1} = i_{n-1}, \dots, N_1 = i_1) \\ &= \frac{\Pr(N_{n+1} = k, N_n = i, N_{n-1} = i_{n-1}, \dots, N_1 = i_1)}{\Pr(N_n = i, N_{n-1} = i_{n-1}, \dots, N_1 = i_1)} \\ &= \frac{\Pr(N_{n+1} - N_n = k - i, N_n - N_{n-1} = i - i_{n-1}, \dots, N_2 - N_1 = i_2 - i_1, N_1 = i_1)}{\Pr(N_n - N_{n-1} = i - i_{n-1}, \dots, N_2 - N_1 = i_2 - i_1, N_1 = i_1)} \\ &= \frac{\Pr(X_{n+1} = k - i, X_n = i - i_{n-1}, \dots, X_1 = i_2 - i_1, N_1 = i_1)}{\Pr(X_n = i - i_{n-1}, X_{n-1} = i_{n-1} - i_{n-2}, \dots, X_1 = i_2 - i_1, N_1 = i_1)} \\ &= \Pr(X_{n+1} = k - i | X_n = i - i_{n-1}, \dots, X_1 = i_2 - i_1, N_1 = i_1) \\ &= \Pr(X_{n+1} = k - i | N_n = i, N_{n-1} = i_{n-1}, \dots, N_1 = i_1) \\ &= \Pr\left(X_{n+1} = k - i \left| \sum_{i=1}^n X_i = i, \sum_{i=1}^{n-1} X_i = i_{n-1}, \dots, N_1 = i_1 \right.\right) \end{aligned}$$

$$\begin{aligned}
&= \Pr(X_{n+1} = k - i) = \Pr(X_{n+1} = k - i) \Pr\left(\sum_{i=1}^n X_i = i\right) / \Pr\left(\sum_{i=1}^n X_i = i\right) \\
&= \Pr\left(X_{n+1} = k - i, \sum_{i=1}^n X_i = i\right) / \Pr\left(\sum_{i=1}^n X_i = i\right) = \Pr\left(X_{n+1} = k - i \mid \sum_{i=1}^n X_i = i\right) \\
&= \Pr(X_{n+1} = k - i \mid N_n = i) = \Pr(N_{n+1} = k \mid N_n = i),
\end{aligned}$$

where the third equality is due to the independence of X_{n+1} and the other n random variables. Thus, the future of a process depends only on the most recent past outcome.

The probability $\Pr(N_{n+1} = k \mid N_n = i, N_{n-1} = i_{n-1}, \dots, N_1 = i_1)$ depends on the value of N_n and is independent of the values of N_1, N_2, \dots, N_{n-1} , since $N_{n+1} = N_n + X_{n+1}$, i.e.,

$$\Pr(N_{n+1} = j \mid N_n = i) = p, \quad \Pr(N_{n+1} = i \mid N_n = i) = 1 - p.$$

So, N_{n+1} depends only on N_n , and both the state space and the parameter set are discrete, then the process $\{N_n, n = 1, 2, \dots\}$ is an example of MC, on the state space $SS = \{0, 1, 2, \dots\}$, with parameter set $T = \{1, 2, \dots\}$ and one-step transition probability:

$$p_{ij} = \Pr(N_{n+1} = j \mid N_n = i) = \Pr(X_{n+1} = j - i) = \begin{cases} p, & j = i + 1, \\ 1 - p, & j = i, \\ 0, & \text{otherwise} \end{cases}$$

$$\text{The TPM is } \mathbf{M} = (p_{ij})_{i,j \in SS} = \begin{matrix} 0 \\ 1 \\ 2 \\ \vdots \end{matrix} \begin{pmatrix} 1-p & p & 0 & \dots \\ 0 & 1-p & p & 0 & \dots \\ 0 & 0 & 1-p & p & \\ \vdots & \vdots & & \ddots & \ddots \\ 0 & \dots & 0 & 0 & \end{pmatrix}.$$

In addition, the MC $\{N_n\}_{n \in \mathbb{N}}$ is time homogeneous if the random sequence $\{X_n\}_{n \geq 1}$ is identically distributed.

Lemma (1) (Probability distribution)

The number N_n of successes in the first n Bernoulli trials, with probability of a success in any one trial is p , is a binomial random variable:

$$\begin{aligned} & \Pr(k \text{ successes in any sequence in } n \text{ trials}) \\ &= \Pr(N_n = k) = \binom{n}{k} p^k q^{n-k}, \text{ for } p + q = 1, k = 0, 1, \dots, n. \quad (\text{i}) \end{aligned}$$

Proof. The state probability at time $n + 1$ can be determined from

the relation: $p_k^{(n+1)} = \sum_i p_i^{(n)} p_{ik}$, i.e.,

$$p_k^{(n+1)} = \Pr(N_{n+1} = k) = \sum_i \Pr(N_n = i) \Pr(N_{n+1} = k | N_n = i),$$

since $\Pr(N_{n+1} = k | N_n = i) = \Pr(X_{n+1} = N_{n+1} - N_n = k - i) = \begin{cases} q, & i = k \\ p, & i = k - 1 \\ 0, & \text{otherwise} \end{cases}$,

the following recurrence relation follows

$$p_k^{(n+1)} = \Pr(N_{n+1} = k) = p \Pr(N_n = k - 1) + q \Pr(N_n = k). \quad (\text{ii})$$

Using the mathematical induction with the recurrence relation (ii), we will prove formula (i):

For $n = 0$, formula (i) is true, since $N_0 = 0$.

Assume that it is also true for $n = m$ and all k . That is

$$\Pr(N_m = k) = \binom{m}{k} p^k q^{m-k}, \text{ for } k = 0, 1, \dots, m. \quad (\text{iii})$$

We will prove it is true for $n = m + 1$, i.e.,

$$\Pr(N_{m+1} = k) = \binom{m+1}{k} p^k q^{m+1-k}. \quad (\text{iv})$$

Putting formula (iii) in the recurrence relation (ii), we get

$$\begin{aligned} \text{L.H.S. } \Pr(N_{m+1} = k) &= p \Pr(N_m = k - 1) + q \Pr(N_m = k) \\ &= p \binom{m}{k-1} p^{k-1} q^{m-k+1} + q \binom{m}{k} p^k q^{m-k} \\ &= \binom{m}{k-1} p^k q^{m-k+1} + \binom{m}{k} p^k q^{m-k+1} = \left[\binom{m}{k-1} + \binom{m}{k} \right] p^k q^{m-k+1} \\ &= \binom{m+1}{k} p^k q^{m-k+1} = \text{R.H.S.}, \text{ for } 0 < k \leq m+1 \end{aligned}$$

when $k = 0$, we get $\Pr(N_{m+1} = 0) = p \times 0 + q \Pr(N_m = 0) = q \times q^m = q^{m+1}$

Therefore, the probability of k successes in n trials is given by a

binomial distribution: $\Pr(N_n = k) = \binom{n}{k} p^k q^{n-k}$, for $k = 0, 1, \dots, n$.

The n^{th} -step transition probabilities

$$\begin{aligned} p_{ij}^{(n)} &= \Pr(N_{m+n} = j \mid N_m = i) = \Pr\left(\sum_{k=1}^{m+n} X_k = j \mid \sum_{k=1}^m X_k = i\right) \\ &= \Pr\left(\sum_{k=m+1}^{m+n} X_k = j - i\right) = \Pr(N_n = j - i) \end{aligned}$$

$$= \begin{cases} \binom{n}{j-i} p^{j-i} q^{n-(j-i)}, & j = i, i+1, \dots, i+n; \quad p+q=1 \\ 0, & j < i \end{cases}$$

The n^{th} -step TPM

$$\mathbf{M}^{(n)} = \left(p_{ij}^{(n)} \right)_{i,j \in SS} = \begin{pmatrix} q^n & npq^{n-1} & \binom{n}{2} p^2 q^{n-2} & \binom{n}{3} p^3 q^{n-3} & \dots & \binom{n}{j} p^j q^{n-j} & \dots \\ 0 & q^n & npq^{n-1} & \binom{n}{2} p^2 q^{n-2} & \ddots & \binom{n}{j-1} p^{j-1} q^{n-(j-1)} & \dots \\ 0 & 0 & q^n & npq^{n-1} & \ddots & & \\ 0 & 0 & 0 & q^n & \ddots & & \vdots \\ 0 & 0 & 0 & 0 & \ddots & & \\ \vdots & 0 & \dots & & 0 & q^n & \\ & & & & \dots & 0 & \ddots \end{pmatrix}.$$

Unlike the Bernoulli process, the mean of the binomial process is dependent on time n and the autocorrelation function is dependent on both m, n and hence is a **non-stationary** process.

Note that the increment $N_{n+m} - N_m$ represents the number of successes through the trials $m+1, m+2, \dots, m+n$:

$$N_{m+n} - N_m = \sum_{k=1}^{m+n} X_k - \sum_{k=1}^m X_k = \sum_{k=m+1}^{m+n} X_k = X_{m+1} + X_{m+2} + \dots + X_{m+n}.$$

It is also a sum of n independent random variables that have the same distribution of the Bernoulli distribution, from which we conclude that

$$\Pr(N_{m+n} - N_m = j) = \Pr(N_n = j) = \binom{n}{j} p^j (1-p)^{n-j}, \quad j = 0, 1, \dots, n.$$

It is the element number j in the binomial expansion $(p+q)^n$ and it does not depend on m where $q=1-p$, for example

Corollary (1). If $n_j > \dots > n_2 > n_1 > n_0 = 0$ are positive integers, then the random variables (increments) $N_{n_j} - N_{n_{j-1}}, \dots, N_{n_2} - N_{n_1}, N_{n_1} - N_{n_0}$ are independent.

Example (1). Find the following

- the JPMF $\Pr(N_{13} = 8, N_7 = 5, N_5 = 4)$
- the expected value $E[N_5 N_8]$.

Solution. Since the following two events are equivalent

$$\{N_{13} = 8, N_7 = 5, N_5 = 4\} \text{ and } \{N_5 = 4, N_7 - N_5 = 1, N_{13} - N_7 = 3\},$$

$$\begin{aligned} \text{then } \Pr(N_{13} = 8, N_7 = 5, N_5 = 4) &= \Pr(N_5 = 4, N_7 - N_5 = 1, N_{13} - N_7 = 3) \\ &= \Pr(N_{13} - N_7 = 3 | N_5 = 4, N_7 - N_5 = 1) \Pr(N_5 = 4, N_7 - N_5 = 1) \\ &= \Pr(N_{13} - N_7 = 3) \Pr(N_5 = 4, N_7 - N_5 = 1) \\ &\quad (\text{since } N_{13} - N_7 \text{ is independent of } N_1, N_2, \dots, N_7) \\ &= \Pr(N_{13} - N_7 = 3) \Pr(N_7 - N_5 = 1 | N_5 = 4) \Pr(N_5 = 4) \\ &= \Pr(N_{13} - N_7 = 3) \Pr(N_7 - N_5 = 1) \Pr(N_5 = 4) \\ &\quad (\text{since } N_7 - N_5 \text{ is independent of } N_1, N_2, \dots, N_5) \\ &= \Pr(N_6 = 3) \Pr(N_2 = 1) \Pr(N_5 = 4) \\ &= \binom{6}{3} p^3 q^3 \binom{2}{1} p q \binom{5}{4} p^4 q = 20 p^8 q^8. \end{aligned}$$

To calculate $E[N_5 N_8]$, we write N_8 as $N_8 = N_5 + (N_8 - N_5)$. Then

$$\begin{aligned} E[N_5 N_8] &= E[N_5 (N_5 + (N_8 - N_5))] = E[N_5^2 + N_5 (N_8 - N_5)] \\ &\quad (\text{since } (N_8 - N_5) \text{ and } N_5 \text{ are independent}) \\ &= E[N_5^2] + E[N_5] E[N_8 - N_5] = E[N_5^2] + E[N_5] E[N_3] \\ &= (5pq + 25p^2) + (5p)(3p) = 5p(q + 8p). \end{aligned}$$

Since each X_i is independent of X_j , $j \neq i$, we conclude that the process N_n is an independent increment process. The independent increments are stationary because

$$\Pr(N_n - N_m = k) = \binom{n-m}{k} p^k (1-p)^{(n-m)-k}, \quad n > m$$

is dependent only on the count difference $(n-m)$ and not on individual counts n and m .

Times of which the Successes of a Bernoulli Process Occur

Denote the times corresponding to the successes in the Bernoulli process by T_1, T_2, T_3, \dots for example if

$$X_1 = 0, X_2 = 1, X_3 = 0, X_4 = 1, X_5 = 1, \dots \text{ then } T_1 = 2, T_2 = 4, T_3 = 5, \dots$$

Relations Between Times and Numbers of Successes

Assume that the success number k has occurred at or before the trial number n , this means that $T_k \leq n$. Then the number of successes in the first n trial should be at least k , meaning that $N_n \geq k$: If $T_k \leq n$ then $N_n \geq k$ and the reverse is true, i.e., if $N_n \geq k$ then $T_k \leq n$.

Assume that $T_n = n$, this achieves the presence $k-1$ of successes in the first of the $n-1$ trials and the success of an event in the trial number n , meaning that $N_n = k-1$ and $X_n = 1$. Conversely, if $N_{n-1} = k-1$ and $X_n = 1$ then $T_n = n$.

We will place the previous two relationships as a corollary, and use them to infer the probability distribution of the time T_n with the knowledge of the probability distribution of N_n .

Corollary (1). For integer numbers $k = 1, 2, \dots$ and $n \geq k$, we have

$$T_k \leq n \text{ iff } N_n \geq k$$

$$T_n = n \text{ iff } N_{n-1} = k - 1 \text{ and } X_n = 1$$

Lemma (3). Let T_n be the time of the n^{th} success in a Bernoulli process $\{X_n, n = 0, 1, \dots\}$. The sequence of random variables $\{T_n, n = 0, 1, \dots\}$ is a MC, with transition probabilities

$$p_{ij} = \Pr(T_n = j | T_{n-1} = i) = \Pr(T_n - T_{n-1} = j - i) = \begin{cases} pq^{j-i-1}, & j \geq i + 1 \\ 0, & \text{otherwise} \end{cases}$$

Here the state space is $SS = \{0, 1, 2, \dots\}$, $T_0 = 0$, and the TPM of $\{T_n, n = 0, 1, 2, \dots\}$ is

$$\mathbf{M} = (p_{ij})_{i,j} = \begin{bmatrix} 0 & p & pq & pq^2 & pq^3 & \dots \\ 0 & 0 & p & pq & pq^2 & \dots \\ 0 & 0 & 0 & p & pq & \dots \\ \vdots & & & 0 & p & \dots \\ 0 & & & & \ddots & \ddots \end{bmatrix}.$$

The state probability at time t is given by

$$p_n^{(t)} = \Pr(T_n = t) = \binom{t-1}{n-1} p^n q^{t-n}, \text{ for } t = n, n+1, \dots$$

with the initial distribution $p_0^{(0)} = 1$; and $p_1^{(0)} = p_2^{(0)} = \dots = 0$.

The n -step transition probabilities of the $\{T_n, n = 0, 1, \dots\}$ MC can be computed as $p_{ij}^{(n)} = \Pr(T_{k+n} = j | T_k = i) = \Pr(T_{k+n} - T_k = j - i)$

$$= \binom{j-i-1}{n-1} p^n q^{j-i-n}, \quad j \geq i + n.$$

The n -step TPM of the $\{T_n, n = 0, 1, \dots\}$ MC is

$$\mathbf{M}^{(n)} = (p_{ij}^{(n)})_{i,j} = \begin{bmatrix} 0 & \dots & 0 & p^n & np^n q & \binom{n+1}{n-1} p^n q^2 & \binom{n+2}{n-1} p^n q^3 & \dots \\ 0 & & 0 & p^n & np^n q & \binom{n+1}{n-1} p^n q^2 & \ddots & \\ \vdots & & & \ddots & p^n & \ddots & \ddots & \end{bmatrix}.$$